

Linear regression

$$\begin{bmatrix} \tilde{x} \\ \tilde{y} \end{bmatrix} \begin{bmatrix} x \\ y \end{bmatrix}$$

Consider

$$t = \underline{\phi}^T(\underline{x}) \underline{w} + \epsilon$$

"t" is a linear function of the vector \underline{w}

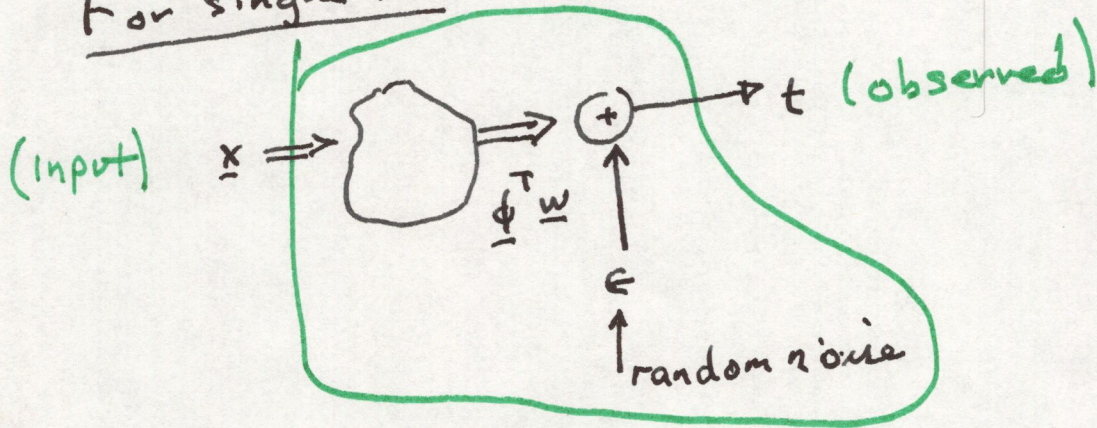
where $\underline{x} = \begin{Bmatrix} x_1 \\ x_2 \\ \vdots \\ x_M \end{Bmatrix}$ $\underline{\phi}(\underline{x}) = \begin{Bmatrix} \phi_1(\underline{x}) \\ \phi_2(\underline{x}) \\ \vdots \\ \phi_M(\underline{x}) \end{Bmatrix} \Leftarrow \text{given}$

the unknown

$$\underline{w} = \begin{Bmatrix} w_1 \\ w_2 \\ \vdots \\ w_{M-1} \\ w_M \end{Bmatrix}$$

N # trial
M # order

For single trial



$$\epsilon: \int_{\epsilon} f(\epsilon)$$

Problem

Given

- ① $\underline{\phi}^T \underline{w}$
- ② $\epsilon: N(0, \sigma^2)$
- ③ $t = \underline{\phi}^T \underline{w} + \epsilon$

find $f_T(t)$

$$\int_{-\infty}^t f_T(z) dz = \int_{-\infty}^{\epsilon} f_{\epsilon}(\beta) d\beta$$

however $\epsilon = t - \underline{\phi}^T \underline{w}$

therefore

$$\int_{-\infty}^t f_T(z) dz = \int_{-\infty}^{t - \underline{\phi}^T \underline{w}} f_{\epsilon}(\beta) d\beta$$

$$\frac{\partial}{\partial t} \left[\int_{-\infty}^t f_T(z) dz \right] = \frac{\partial}{\partial t} \left[\int_{-\infty}^{t - \underline{\phi}^T \underline{w}} f_{\epsilon}(\beta) d\beta \right]$$

$$f_T(t) = f_{\epsilon}(t - \underline{\phi}^T \underline{w})$$

Recall

$$\epsilon: N(0, \sigma^2)$$

hence

$$f_{\epsilon}(\epsilon) = \frac{e^{-\left[\frac{\epsilon^2}{2\sigma^2}\right]}}{\sqrt{2\pi\sigma^2}}$$

The pdf of t is

$$f_T(t) = \frac{e^{-\frac{(t - \underline{\phi}^T \underline{\omega})^2}{2\sigma^2}}}{\sqrt{2\pi\sigma^2}}$$

$$\hat{N}(t | \underline{\phi}^T \underline{\omega}) \equiv f_T(t)$$

$$E(t) = \underline{\phi}^T \underline{\omega}$$

$$E(t^2) - E(t)^2 = \sigma^2.$$

N-trials (~~independ~~ independent)

$$\underline{t} = \{t_1, t_2, \dots, t_N\}$$

$$\underline{X} = \{x_1, x_2, \dots, x_N\}$$

$$P(\underline{t} | \underline{X}, \underline{w}) = \prod_{n=1}^N \hat{N}(t_n | \underline{\phi}^T(x_n) \underline{w})$$

$$= \prod_{n=1}^N \frac{e^{-\frac{(t_n - \underline{\phi}^T(x_n) \underline{w})^2}{2\sigma^2}}}{\sqrt{2\pi\sigma^2}}$$

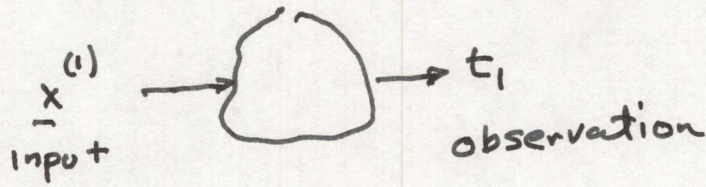
$$P(\underline{t} | \underline{X}, \underline{w}) = \frac{e^{-\sum_{n=1}^N \frac{(t_n - \underline{\phi}^T(x_n) \underline{w})^2}{2\sigma^2}}}{(\sqrt{2\pi\sigma^2})^N}$$

Simplified

$$[\Phi] = \begin{bmatrix} \phi_1(x_1) & \dots & \phi_m(x_1) \\ \vdots & & \vdots \\ \phi_1(x_N) & \dots & \phi_m(x_N) \end{bmatrix}$$

$$[\Phi] = \begin{bmatrix} \underline{\phi}^T(x_1) \\ \underline{\phi}^T(x_2) \\ \vdots \\ \underline{\phi}^T(x_N) \end{bmatrix}$$

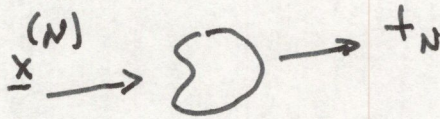
Example



trial 1



trial 2



trial N

Single trial

Given $\underline{x}^{(j)} = \begin{bmatrix} 1 \\ x_2^{(j)} \\ x_3^{(j)} \end{bmatrix} \quad \& \quad t_j$

Model

$$t_j = \underline{\phi}^T(\underline{x}^{(j)}) \underline{w} + \epsilon^{(j)}$$

define/choose Basis

$$\underline{\phi}(\underline{x}^{(j)}) = \begin{bmatrix} \phi_1(\underline{x}^{(j)}) \\ \phi_2(\underline{x}^{(j)}) \\ \phi_3(\underline{x}^{(j)}) \end{bmatrix} = \begin{bmatrix} 1 \\ x_2^{(j)} \\ x_3^{(j)} \end{bmatrix}$$

Unknown weights \underline{w}

$$\underline{w} = \begin{bmatrix} w_1 \\ w_2 \\ w_3 \end{bmatrix}$$

The model becomes

$$t_j = \begin{bmatrix} 1 & x_2^{(j)} & x_3^{(j)} \end{bmatrix} \begin{bmatrix} w_1 \\ w_2 \\ w_3 \end{bmatrix} + e^{(j)}$$

$$t_j = w_1 + x_2^{(j)} w_2 + x_3^{(j)} w_3 + e^{(j)}$$

Combining the N observations

$$\begin{bmatrix} t_1 \\ t_2 \\ t_3 \\ \vdots \\ t_N \end{bmatrix} = \begin{bmatrix} 1 & x_1^{(1)} & x_2^{(1)} \\ 1 & x_1^{(2)} & x_2^{(2)} \\ 1 & x_1^{(3)} & x_2^{(3)} \\ \vdots & \vdots & \vdots \\ 1 & x_1^{(N)} & x_2^{(N)} \end{bmatrix} \begin{bmatrix} w_1 \\ w_2 \\ w_3 \end{bmatrix} + \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \\ \vdots \\ \epsilon_N \end{bmatrix}$$

$$\underline{t} = \Phi \underline{w} + \underline{\epsilon}$$

$$-\frac{[\underline{t} - \Phi \underline{w}]^T [\underline{t} - \Phi \underline{w}]}{2\sigma^2}$$

$\| \underline{t} - \Phi \underline{w} \|^2$

$$P(\underline{t} | \underline{X}, \underline{w}) = \frac{e}{(2\pi\sigma^2)^{N/2}}$$

Log likelihood

$$\ln [P(\underline{t} | \underline{X}, \underline{w})]$$

find \underline{w} that maximized the
log likelihood of \underline{t}

Maximize

$$\ln(p(\underline{t} | \underline{X}, \underline{w})) =$$

$$= -\frac{N}{2} \ln[2\pi] - \frac{N}{2} \ln(\sigma^2) - \frac{[\underline{t} - \Phi \underline{w}]^T [\underline{t} - \Phi \underline{w}]}{2\sigma^2}$$

or minimize the $-\ln(p(\underline{t} | \underline{X}, \underline{w})) \equiv F$

$$F = \frac{N}{2} \ln[2\pi] + \frac{N}{2} \ln(\sigma^2) + \frac{[\underline{t} - \Phi \underline{w}]^T [\underline{t} - \Phi \underline{w}]}{2\sigma^2}$$

$$1^{\circ} \frac{\partial F}{\partial \underline{w}} = 0 \Rightarrow \Phi^T [\underline{t} - \Phi \underline{w}] = 0$$

solu:

$$\underline{w}_{MLE} = [\Phi^T \Phi]^{-1} \Phi^T \underline{t}$$

$$2^{\circ} \frac{\partial F}{\partial \sigma^2} = \frac{N}{2} \frac{1}{\sigma^2} + [\underline{t} - \Phi \underline{w}]^T [\underline{t} - \Phi \underline{w}] \left\{ \frac{-1}{2(\sigma^2)^2} \right\} = 0$$

$$\sigma^2 = \frac{1}{N} [\underline{t} - \Phi \underline{w}]^T [\underline{t} - \Phi \underline{w}] \Big|_{\underline{w} = \underline{w}_{MLE}}$$

MAD \rightarrow $P(\underline{w} | \underline{t})$
MLE \rightarrow $P(\underline{t} | \underline{w})$

Proof \underline{w}_{ml}

$$\begin{aligned}\frac{\partial F}{\partial \underline{w}} &= \frac{1}{2\sigma^2} \frac{\partial}{\partial \underline{w}} \left[(\underline{t}^T - \underline{w}^T \phi^T) (\underline{t} - \phi \underline{w}) \right] \\ &= \frac{1}{2\sigma^2} \frac{\partial}{\partial \underline{w}} \left[\underline{t}^T \underline{t} - \underline{t}^T \phi \underline{w} - \underline{w}^T \phi^T \underline{t} + \underline{w}^T \phi^T \phi \underline{w} \right]\end{aligned}$$

note $\frac{\partial}{\partial \underline{a}} (\underline{a}^T \phi \underline{b}) = \phi \underline{b}$

$\frac{\partial}{\partial \underline{b}} (\underline{a}^T \phi \underline{b}) = (\underline{a}^T \phi)^T = \phi^T \underline{a}$

$$= \frac{1}{2\sigma^2} \left[-\phi^T \underline{t} - \phi^T \underline{t} + \phi^T \phi \underline{w} + \phi^T \phi \underline{w} \right]$$

$$= \frac{1}{2\sigma^2} \left[-2 \cdot \phi^T \underline{t} + 2 \phi^T \phi \underline{w} \right] = 0$$

$$\phi^T \phi \underline{w} = \phi^T \underline{t}$$

$$\underline{w} = (\phi^T \phi)^{-1} \phi^T \underline{t}$$

Max a Posterior prob (MAP)

Bayes thm

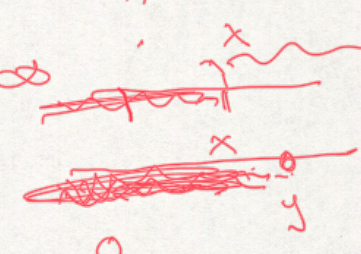
$$P(\underline{w} | \underline{t}) = \frac{P(\underline{t} | \underline{w}) P(\underline{w})}{P(\underline{t})}$$

1^o prior $P(\underline{w}) = N_{\underline{w}}(0, I \alpha^2)$

2^o Likelihood

$$P(\underline{t} | \underline{w}) = \prod_{n=1}^N \hat{N}(t_n | \phi^T \underline{w})$$

$$P(\underline{w} | \underline{t}) = \frac{e^{-\frac{\|\underline{t} - \Phi \underline{w}\|^2}{2\sigma^2}}}{(2\pi\sigma^2)^{N/2}} \frac{e^{-\frac{\underline{w}^T \underline{w}}{2\alpha^2}}}{(2\pi\alpha^2)^{M/2}} \frac{1}{P(\underline{t})}$$



$$y = ax + \epsilon \quad a < 0$$

$$\text{Prob}(\underline{X} \leq y) = \text{Prob}(a\underline{X} + \epsilon \leq y)$$

$$\text{Prob}(a\underline{X} \leq y - \epsilon)$$

$$\text{Prob}(\underline{X} \leq \frac{y - \epsilon}{a})$$

$$\text{Prob}(\underline{X} > \frac{y - \epsilon}{a})$$

$$1 - \text{Prob}(\underline{X} < \frac{y - \epsilon}{a})$$

$$y = ax + \epsilon$$

$$\frac{y - \epsilon}{a} = x$$

MAP (estimate)

$$\ln[P(\underline{w} | \underline{t})] = \ln(P(\underline{t} | \underline{w})) + \ln(P(\underline{w})) - \ln(P(\underline{t}))$$

$$F = -\ln(P(\underline{t})) - \frac{N}{2} \ln(2\pi) - \frac{N}{2} \ln(\sigma^2) - \frac{M}{2} \ln(\alpha^2) - \frac{M}{2} \ln(2\pi) \\ - \frac{\|\underline{t} - \Phi \underline{w}\|^2}{2\sigma^2} - \frac{\underline{w}^T \underline{w}}{2\alpha^2}$$

maximize F

$$\frac{\partial F}{\partial \underline{w}} = 0 = - \left[\frac{\Phi^T \underline{t} - \Phi^T \Phi \underline{w}}{\sigma^2} \right] - \frac{\underline{w}}{\alpha^2} = 0$$

$$\left[\frac{\Phi^T \Phi}{\sigma^2} + \frac{\underline{I}}{\alpha^2} \right] \underline{w} = \frac{\Phi^T \underline{t}}{\sigma^2}$$

$$\frac{\partial F}{\partial \sigma^2} = 0 = -\frac{N}{2} \frac{1}{\sigma^2} + \frac{\|\underline{t} - \Phi \underline{w}\|^2}{2(\sigma^2)^2} = 0$$

$$\sigma^2 = \frac{1}{N} \|\underline{t} - \Phi \underline{w}\|^2$$

$$\underline{w}_{\text{MAP}} = \left[\Phi^T \Phi + \frac{\underline{I}}{\alpha^2} \right]^{-1} \left[\Phi^T \underline{t} \right]$$

\downarrow
 $(\underline{I} \beta^2)$

Parametric distribution
of \underline{w} $\{P(\underline{w} | \underline{t})\}$

\underline{w} | \downarrow - MLE
- abs MAP

$$P(\underline{w} | \underline{t}) P(\underline{t}) = P(\underline{t} | \underline{w}) P(\underline{w})$$

① prior $P(\underline{w})$

$$P(\underline{w}) = \frac{e^{-[\underline{w} - \underline{m}_0]^T \Sigma_0^{-1} [\underline{w} - \underline{m}_0] \frac{1}{2}}}{(2\pi)^{M/2} |\Sigma_0|^{1/2}}$$

where $\underline{w} = \begin{Bmatrix} w_1 \\ w_2 \\ \vdots \\ w_M \end{Bmatrix}$; \underline{m}_0 mean; Σ_0 covariance of \underline{w}

② likely hood

$$P(\underline{t} | \underline{w}) = \frac{e^{-\frac{\|\underline{t} - \Phi \underline{w}\|^2}{2\sigma^2}}}{(2\pi\sigma^2)^{N/2}}$$

$$- [\underline{w} - \underline{m}_N]^T \Sigma_N^{-1} [\underline{w} - \underline{m}_N] \frac{1}{2}$$

③ $P(\underline{w} | \underline{t}) P(\underline{t}) \propto \frac{e^{-[\underline{w} - \underline{m}_N]^T \Sigma_N^{-1} [\underline{w} - \underline{m}_N] \frac{1}{2}}}{(2\pi)^{M/2} |\Sigma_N|^{1/2}}$

estimate $\underline{m}_N, \Sigma_N$

model

$$\underline{m}_N, \Sigma_N$$

Moment matching to find

$$\underline{m}_N \in \Sigma_N$$

$$\frac{1}{2} [\underline{w} - \underline{m}_0]^T \Sigma_0^{-1} [\underline{w} - \underline{m}_0] + [\underline{t} - \Phi \underline{w}]^T [\underline{t} - \Phi \underline{w}] \frac{1}{2\sigma^2}$$

$$= [\underline{w} - \underline{m}_N]^T \Sigma_N^{-1} [\underline{w} - \underline{m}_N] \frac{1}{2}$$

$$\frac{1}{\sigma^2} \left[\underline{w}^T \Phi^T \Phi \underline{w} + \underline{t}^T \underline{t} - \underline{w}^T \Phi^T \underline{t} - \underline{t}^T \Phi \underline{w} \right]$$

$$+ \left[\underline{w}^T \Sigma_0^{-1} \underline{w} + \underline{m}_0^T \Sigma_0^{-1} \underline{m}_0 - \underline{w}^T \Sigma_0^{-1} \underline{m}_0 - \underline{m}_0^T \Sigma_0^{-1} \underline{w} \right]$$

$$= \left[\underline{w}^T \Sigma_N^{-1} \underline{w} + \underline{m}_N^T \Sigma_N^{-1} \underline{m}_N - \underline{w}^T \Sigma_N^{-1} \underline{m}_N - \underline{m}_N^T \Sigma_N^{-1} \underline{w} \right]$$

$$\textcircled{1} \quad \underline{w}^T \frac{\Phi^T \Phi}{\sigma^2} \underline{w} + \underline{w}^T \Sigma_0^{-1} \underline{w} = \underline{w}^T \Sigma_N^{-1} \underline{w}$$

$$\boxed{\frac{\Phi^T \Phi}{\sigma^2} + \Sigma_0^{-1} = \Sigma_N^{-1}}$$

$$\textcircled{2} \quad \underline{t}^T \frac{\Phi}{\sigma^2} \underline{w} + \underline{m}_0^T \Sigma_0^{-1} \underline{w} = \underline{m}_N^T \Sigma_N^{-1} \underline{w}$$

$$\underline{t}^T \Phi + \underline{m}_0^T \Sigma_0^{-1} = \underline{m}_N^T \Sigma_N^{-1}$$

$$\left(\frac{t^T \phi}{\sigma^2} + m_0^T \Sigma_0^{-1} \right)^T = \left(m_N^T \Sigma_N^{-1} \right)^T$$

note $(\Sigma_N^{-1})^T = \Sigma_N^{-1}$ Σ_N^{-1} is symmetric
 likewise Σ_0^{-1} is sym

$$\left[\frac{\phi^T t}{\sigma^2} + \Sigma_0^{-1} m_0 \right] = \Sigma_N^{-1} m_N$$

$$m_N = \Sigma_N \left[\frac{\phi^T t}{\sigma^2} + \Sigma_0^{-1} m_0 \right]$$